

## Lecture Notes 12

### 29 Pseudorandom Generators

**Definition:** An ensemble  $X = \{X_n\}_{n \in \mathbb{N}}$  is *pseudorandom* if  $X, U$  are indistinguishable in polynomial time, where  $U = \{U_n\}_{n \in \mathbb{N}}$  is the uniform ensemble.

Thus,  $X$  is pseudorandom if it “looks” the same to all probabilistic polynomial time algorithms.

**Definition:** A *pseudorandom generator* is a deterministic polynomial time function  $G$  that satisfies two properties:

1.  $G$  maps strings of length  $n$  to strings of length  $\ell(n) > n$ .  $\ell(n)$  is called the *expansion factor*.
2.  $\{G(U_n)\}_{n \in \mathbb{N}}$  is pseudorandom.

We remark that if  $G$  is a pseudorandom generator, then  $G(U_n)$  is not statistically close to  $U_{\ell(n)}$ . To see this, let  $R_G = \{G(x) \mid x \in \{0, 1\}^n\}$  be the range of  $G$ . Clearly,  $|R_G| \leq 2^n$ , and for all  $y \notin R_G$ ,  $\Pr[G(U_n) = y] = 0$ . On the other hand, for the uniform ensemble,  $\Pr[U_{\ell(n)} = y] = \frac{1}{2^{\ell}}$ . Hence, the statistical difference

$$\begin{aligned} \Delta(\ell(n)) &= \frac{1}{2} \sum_{\alpha \in \{0,1\}^{\ell(n)}} |\Pr[G(U_n) = \alpha] - \Pr[U_{\ell(n)} = \alpha]| \\ &\geq \frac{1}{2} \sum_{\alpha \in R_G} |\Pr[G(U_n) = \alpha] - \Pr[U_{\ell(n)} = \alpha]| \\ &= \frac{1}{2} \sum_{\alpha \in R_G} \left|0 - \frac{1}{2^{\ell}}\right| \\ &= \frac{1}{2} \cdot \frac{2^{\ell} - 2^n}{2^{\ell}} \geq \frac{1}{4}, \end{aligned}$$

is not negligible, so  $G(U_n)$  and  $U_{\ell(n)}$  are not statistically close.

We now describe how to build a pseudorandom number generator  $G$  with polynomial expansion factor starting from a generator  $G_1$  with expansion factor  $\ell(n) = n + 1$ .

Fix a polynomial  $p(n)$ . For  $s \in \{0, 1\}^n$ , write the length- $(n + 1)$  string  $G_1(s)$  as  $\sigma s'$ , where  $|\sigma| = 1$  and  $|s'| = n$ . On input  $s$ , iteratively define the sequences  $s_0, s_1, s_2, \dots, s_{p(n)}$  and  $\sigma_1, \sigma_2, \dots, \sigma_{p(n)}$  as follows:

$$\begin{aligned} s_0 &= s \\ \sigma_i s_i &= G_1(s_{i-1}), \text{ for } i = 0, 1, 2, \dots, p(n) - 1. \end{aligned}$$

The output of  $G(s)$  is the sequence  $\sigma_1 \sigma_2 \dots \sigma_{p(n)}$ .  $G(s)$  is easily computed in polynomial time by a simple iterative program that calls  $G_1$  a total of  $p(n)$  times.

**Theorem 1** *If  $G_1$  is pseudorandom, then so is  $G$ .*

Proof is by a hybrid argument. We let hybrid  $H_n^k$  consist of  $k$  uniform random bits followed by the first  $p(n) - k$  bits of  $G(s_0)$ , which we write as  $G(s_0): [1, p(n) - k]$ . In symbols,

$$H_n^k = U_k \cdot G(U_n): [1, p(n) - k].$$

Clearly,  $H_n^0 = G(U_n)$  and  $H_n^{p(n)} = U_{p(n)}$ .

Suppose  $D$  distinguishes  $G(U_n)$  from  $U_{p(n)}$  with absolute probability difference  $\delta(n)$ . Then for some  $k$ ,  $D$  distinguishes  $H_n^k$  from  $H_n^{k+1}$  with absolute probability difference  $\geq \delta(n)/p(n)$ .

We now describe an algorithm  $D'$  that attempts to distinguish  $G_1(U_n)$  from  $U_{n+1}$ . On length- $(n + 1)$  input  $\alpha$ ,  $D'$  does the following:

1. Write  $\alpha = \tau \cdot \alpha'$ , where  $|\tau| = 1$  and  $|\alpha'| = n$ .
2. Choose index  $k$  uniformly from  $\{0, 1, \dots, p(n) - 1\}$ .
3. Choose a uniformly distributed string  $\beta$  of length  $k$ .
4. Construct  $y = \beta \cdot \tau \cdot G(\alpha'): [1, p(n) - k - 1]$ .
5. Compute and output  $D(y)$ .

If  $\alpha$  is uniformly distributed, then  $\tau$  and  $\alpha'$  are both uniformly distributed, so  $y = H_n^{k+1}$ . On the other hand, if  $\alpha = G_1(s_0)$ , where  $s_0$  is uniformly distributed, then  $\tau = \sigma_1$  and  $\alpha' = s_1$ , so  $y = H_n^k$ . This is because

$$G(s_0): [1, p(n) - k] = \tau \cdot G(s_1): [1, p(n) - k - 1]$$

Hence,  $D'$  distinguishes  $G_1(U_n)$  from  $U_{n+1}$  with absolute probability difference  $\geq \delta(n)/p(n)$ .

We omit the remaining details of showing how this leads to a contradiction of the assumption that  $G$  is not pseudorandom.

## 30 Unpredictability

Our formal definition of pseudorandomness is based on the indistinguishability of an entire polynomial-length generated sequence from a uniformly distributed random sequence. However, the traditional notion of a pseudorandom generator is based on repeated experiments. The output bits  $x_1, x_2, \dots$  are assumed to be generated one at a time. The generator is called pseudorandom if each  $x_i$  “appears” to result from an independent and uniformly distributed random event such as the flip of a fair coin.

The notion of “appears” is can be captured in terms of unpredictability. We say that  $x_{i+1}$  is unpredictable if no polynomial time algorithm that attempts to guess it is correct with more than a tiny advantage over chance, even given all of the prior bits  $x_1, \dots, x_i$ .

More formally, a *predictor* is a p.p.t. algorithm  $A$  that is allowed to read the input sequence  $x$  a bit at a time in order. After reading bit  $i$ , the algorithm can choose to output a guess  $b$  and halt, or it can continue. In any case, it must halt and emit a guess after reading the next-to-last bit of  $x$ . Let  $k$  be the last bit read by  $A$ . Then  $A$  is *correct* if  $k < |x|$  and  $b = x_{k+1}$ . In addition to the input  $x$ , which  $A$  is allowed to read only a bit at a time,  $A$  is also given an input  $1^n$ , where  $n = |x|$ . This way,  $A$  can determine the length of  $x$  without having to read it all.

Notation: The textbook uses the notation  $\text{next}_A(x)$  to denote the next bit of  $x$  following the last bit that  $A$  read. The intent is that the event  $[A(1^{|X_n|}, X_n) = \text{next}_A(X_n)]$  should mean that a string  $x$  is chosen according to the distribution  $X_n$ ,  $A$  is run on inputs  $1^n$  and  $x$ ,  $A$  reads the first  $k$  bits of  $x$  for some  $k$  and outputs  $b$ , and  $b = x_{k+1}$ , the “next” bit of  $x$ . That is, the event is that  $A$  correctly predicts some bit of a randomly chosen  $x$  from distribution  $X_n$ .

A better notation would make  $k$  explicit. For example, we could pretend that  $A$  outputs a pair  $(k, b)$  with the meaning that  $k$  is the index of the last bit of  $x$  that  $A$  read, and  $b$  is  $A$ 's prediction for  $x_{k+1}$ . We could then define  $\text{next}_A(x) = \{(k, x_{k+1}) \mid k \in [0, n-1]\}$ . Now,  $A$  correctly predicts the next bit if  $A(1^{|x|}, x) = (k, b)$  and  $(k, b) \in \text{next}_A(x)$ .

**Definition:** An ensemble  $\{X_n\}_{n \in \mathbb{N}}$  is called *unpredictable in polynomial time* if for every p.p.t.  $A$ , every positive polynomial  $p(\cdot)$ , and all sufficiently large  $n$ ,

$$\Pr[A(1^{|x|}, x) \in \text{next}_A(x)] < \frac{1}{2} + \frac{1}{p(n)}.$$

**Theorem 2** *An ensemble  $X$  is pseudorandom if and only if it is unpredictable in polynomial time.*

**Proof:**

( $\Rightarrow$ ) The theorem in the forward direction is straightforward. We sketch the general ideas and leave the details to the reader.

If there were a predictor  $A$  for  $X$ , then a distinguisher  $D$  is easily built. Namely,  $D(x)$  outputs 1 iff  $A(1^{|x|}, x)$  correctly predicts the next bit. If  $x$  comes from  $X$ ,  $D(x)$  will output 1 with probability at least  $\frac{1}{2} + \frac{1}{p(n)}$ , but if  $x$  comes from  $U$ , then clearly  $D(x)$  will output 1 with probability exactly  $\frac{1}{2}$ . Hence,  $D$  successfully distinguishes  $X$  from  $U$ .

( $\Leftarrow$ ) The theorem in the reverse direction is proved by another hybrid argument. We sketch a few of the main ideas. Assume  $X$  is both unpredictable but not pseudorandom. Then there is a distinguisher  $D$  such that

$$|\Pr[D(X_n) = 1] - \Pr[D(U_n) = 1]| \geq \frac{1}{p(n)}$$

for infinitely many  $n$ . We may without loss of generality drop the absolute value brackets and assume that

$$\Pr[D(X_n) = 1] - \Pr[D(U_n) = 1] \geq \frac{1}{p(n)}$$

for infinitely many  $n$ . The reasoning is that either  $\Pr[D(X_n) = 1] \geq \Pr[D(U_n) = 1]$  for infinitely many  $n$ , or  $\Pr[D(X_n) = 1] \leq \Pr[D(U_n) = 1]$  for infinitely many  $n$ . If the latter, then  $\Pr[\bar{D}(X_n) = 1] \geq \Pr[\bar{D}(U_n) = 1]$  for the algorithm  $\bar{D}$  that is identical to  $D$  except that it complements the output.

We build a next-bit predictor  $A$ . Let hybrid  $H_n^k$  consist of the first  $k$  bits from  $X_n$  followed by the last  $n - k$  bits from  $U_n$ . Then  $H_n^n = X_n$  and  $H_n^0 = U_n$ . The predictor  $A(1^{|x|}, x)$  guesses a number  $k \in [0, |x| - 1]$ , reads only the first  $k$  bits of  $x$ , and constructs the string  $y = x_1, \dots, x_k, u_{k+1}, \dots, u_n$ , where the  $u_j$ 's are uniformly distributed random bits. It then runs  $D(y)$ . If  $D(y) = 1$ , then  $A$  predicts bit  $k + 1$  to be  $u_{k+1}$ . Otherwise,  $A$  predicts bit  $k + 1$  to be  $\neg u_{k+1}$  (the complement of  $u_{k+1}$ ).

We omit the non-trivial analysis needed to show that algorithm  $A$  has a sufficient advantage as a next-bit predictor to contradict the assumption that  $X$  is unpredictable.  $\blacksquare$

## 31 Pseudorandom Generators and One-Way Functions

We now show that the existence of pseudorandom generators implies the existence of one-way functions.

**Theorem 3** Let  $G$  be a pseudorandom generator with expansion factor  $\ell(n) = 2n$ . Define the function  $f(x, y) = G(x)$  for all  $|x| = |y|$ . Then  $f$  is a strongly one-way function.

**Proof:** Suppose  $f$  is not strongly one-way. Let  $A$  be an inverter for  $f(U_{2n})$  with success probability at least  $\frac{1}{p(n)}$  for infinitely many  $n$ . We construct a distinguisher  $D$  that distinguishes  $G(U_n)$  from  $U_{2n}$  on those same  $n$ .

$D(\alpha)$  uses  $A$  to attempt to find  $\beta$  such that  $f(\beta) = \alpha$ . If  $A$  succeeds, then  $D$  outputs 1; otherwise  $D$  outputs 0. Since  $f(U_{2n}) = G(U_n)$ , then

$$\Pr[D(G(U_n)) = 1] = \Pr[f(A(f(U_{2n}))) = f(U_{2n})] \geq \frac{1}{p(n)}. \quad (1)$$

On the other hand,

$$\Pr[D(U_{2n}) = 1] = \Pr[f(A(U_{2n})) = U_{2n}] \leq \frac{1}{2^n}. \quad (2)$$

This is because  $f(x, y)$  depends only on  $x$ , so the range of  $f$  on pairs of length- $n$  inputs has size  $\leq 2^n$ . Since  $f(A(U_{2n}))$  is in the range of  $f$ , the probability that  $U_{2n}$  is in the range, much less actually equal to  $f(A(U_{2n}))$ , is at most  $2^{-n}$ . Subtracting 2 from 1 gives

$$\Pr[D(G(U_n)) = 1] - \Pr[D(U_{2n}) = 1] \geq \frac{1}{p(n)} - \frac{1}{2^n} \geq \frac{1}{2p(n)}. \quad (3)$$

Thus,  $D$  distinguishes  $G(U_n)$  from  $U_{2n}$  for infinitely many  $n$ , contradicting the assumption that  $G$  is a pseudorandom generator. ■